

# OUTAGE PERFORMANCE OF DUAL DIVERSITY RECEIVER OVER Rician FADING CHANNEL

Suryakant Pathak<sup>1</sup>, Himanshu Katiyar<sup>2</sup>

<sup>1</sup>Associate Professor & Head, Deptt of CSE, Dr. K N Modi University Newai, Rajasthan, (India)

& Research scholar, Deptt. of CSE, BBD University, Lucknow, (India)

<sup>2</sup>Associate Professor, Deptt. of ECE, BBDNIIT, Lucknow, (India)

## ABSTRACT

In wireless communication fading is a phenomenon which degrades the performance of the link which includes Rayleigh, weibull, Nakagami-q, Nakagami-m, Rician fading. In this paper the outage performance of dual diversity receiver over Rician fading channel have been analysed. For this purpose probability density function (PDF) of SNR, cumulative distribution function (CDF) and outage performance ( $P_{out}$ ) using different diversity combining techniques selection combining, maximal ratio combining, equal gain combining and switch & stay combining for dual receiver is analysed for different value of Rice factor  $K$ .

**Keywords:** Diversity Combining Techniques, Dual Diversity, Outage Probability, Probability Distribution Function, Rician fading channel.

## I INTRODUCTION

In wireless communication fading is a phenomenon which degrades the performance of the link it includes Rayleigh, weibull, Nakagami-q, Nakagami-m, Rician fading. The Rician fading distribution is often used to model propagation paths, consisting of one strong direct line-of-sight (LoS) signal and many randomly reflected and usually weaker signals. Independent fading path can be achieved by using multiple transmit or receive antenna array, where the elements of array are separated in distance. This type of diversity is called space diversity [1]. Space diversity reception is a very efficient technique for mitigating fading and co-channel interference (CCI) effects to improve the quality of service (QoS) in wireless communication systems. Various diversity combining techniques are used for reducing the fading effects and the influence of the co-channel interference (CCI) in wireless communication systems for which all the calculations are evaluated [2]. In this paper mainly four types of combining techniques selection combining (SC), maximal ratio combining (MRC), equal gain combining (EGC) and switch and stay combining (SSC) have been analysed in respect to their PDF of SNR as well as cumulative distribution function (CDF) or outage probability ( $P_{out}$ ) [3], [4], [5]. Outage performance of dual diversity under Rayleigh fading channels have been studied in [6]. Performance of dual diversity under  $\alpha$ - $\mu$  fading channels has been studied in [7]. Performance selection combining receivers been studied in [8]. Performance of switch and stay combining receivers have been studied in [9], [11]. Performances of EGC and MRC have been studied in [10].

The rest of the paper is organised as follows. In section 2, Rician fading model is briefly discussed. In section 3, diversity combining techniques for Rician fading channel is discussed. In section 4, the simulation result has been analysed, and section 5, concludes the paper.

## II RICIAN FADING MODEL

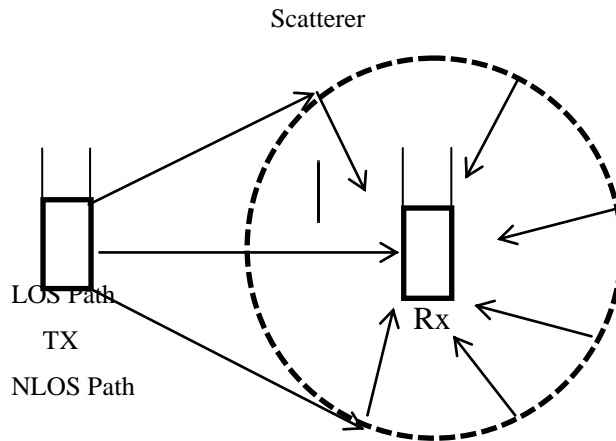


Fig 1. Rician fading model

### 2.1 Rician Fading

Some types of scattering environment have specular or LOS components. If  $g_I(t)$  and  $g_Q(t)$  are Gaussian random process with non-zero mean  $m_I(t)$  and  $m_Q(t)$ . If we again assume that these process are uncorrelated and random variable  $g_I(t)$  and  $g_Q(t)$  have the same variance  $\sigma^2$ . Then magnitude of the received complex envelop at time  $t$  has a Rician distribution as [3, Eq. (2.45)]

$$f_{a_l}(x) = \frac{x}{\sigma^2} e^{-\frac{(x^2+s^2)}{2\sigma^2}} I_0\left(\frac{xs}{\sigma^2}\right) \quad (1)$$

Where

$$l=1, 2$$

$$s^2 = m_I^2(t) + m_Q^2(t) \quad (2)$$

is called the non-centrality parameter. This type of fading is called Ricean fading and is very often observed in microcellular and mobile satellite applications. A very simple Ricean fading model assumes that the means  $m_I(t)$  and  $m_Q(t)$  are constants, i.e.  $m_I(t) = m_I$ , and  $m_Q(t) = m_Q$ . The means  $m_I(t)$  and  $m_Q(t)$  corresponding to the in phase and quadrature components of the LoS signal are given by

$$m_I(t) = s \cdot \cos(2\pi f_m \cos\theta_0 t + \phi_0) \quad (3)$$

$$m_Q(t) = s \cdot \sin(2\pi f_m \sin\theta_0 t + \phi_0) \quad (4)$$

Where  $f_m \cos\theta_0$  and  $\phi_0$  are the Doppler shift and random phase offset associated with the LoS or specular component, respectively.

The Rice factor,  $K$ , is defined as the ratio of the specular power  $s^2$  to scattered power  $2\sigma^2$ , i.e.  $K = s^2/2\sigma^2$ . When  $K=0$  the channel exhibits Rayleigh fading, and when  $K=\infty$  the channel does not exhibit any fading at all. The envelope distribution can be rewritten in terms of the Rice factor and the average envelope power  $E[a^2] = \Omega_p = s^2/2\sigma^2$  by first noting that

$$s^2 = \frac{K\Omega_p}{K+1}, \quad 2\sigma^2 = \frac{\Omega_p}{K+1} \quad (5)$$

It then follows

$$f_{\alpha_l}(x) = \frac{x}{\sigma^2} = \frac{2x(K+1)}{\Omega_p} \exp\left\{-K - \frac{(K+1)x^2}{\Omega_p}\right\} I_0\left(2x\sqrt{\frac{K(K+1)}{\Omega_p}}\right), \quad x > 0 \quad (6)$$

Where

$\Omega_p$  is total received envelop power

$$\Omega_p = E[g_I^2(t)] + E[g_Q^2(t)] = \sum_{n=1}^N c_n^2$$

$c_n$  is the magnitude associated with n preoperational path

$\alpha$  = received complex envelop which is Rician distributed

$$\alpha = N(m_I, \sigma^2) + jN(m_Q, \sigma^2)$$

$N = (.,.)$  Normally or Gaussian distributed random variable

$x$  = running variable

$\sigma^2$  = Variance

$$s = \sqrt{m_I^2 + m_Q^2}$$

$m_I$  = mean of in phase component

$m_Q$  = mean of quadrature phase components

$I_0$  = zero order modified Bessel function of the first kind

$j$  = complex operator

### III DIVERSITY COMBINING TECHNIQUE

Diversity technique provides multiple copies of the same signal on different branches, which undergo independent fading. If one branch undergoes a deep fade, another branch may have strong signal. In space diversity fading is minimized by the simultaneous use of two or more physically separated antennas. Thus having more than one path to select the SNR at receiver may be improved by selecting appropriate combining technique. SNR  $\gamma$  is a random variable and is given for different diversity schemes.

$$\gamma \in \{\gamma_{SC}, \gamma_{MRC}, \gamma_{EGC}, \gamma_{SSC}\}$$

Following diversity combining techniques for Rician fading channels are discussed.

#### 3.1 Selection Combining (SC)

Selection combining is based on the principle of selecting the best signal among all the signals received from different branches at the receiving end. In this method, the receiver monitors the SNR of the incoming signal using switch logic. The branch with highest instantaneous SNR is connected to demodulator. SNR of selection combining is given as

$$\gamma_{SC} = \max(R_1^2, R_2^2)$$

Where  $R_1$  and  $R_2$  represent the fading envelope for two channels seen by two different antennas.

CDF of selection combining with two receive antennas can be calculated from [12, Eq.(11)]

$$F_{\gamma_{SC}}(\gamma) = \left[ \frac{e^{-\frac{s^2}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{\Gamma(k+1, \frac{\gamma}{\gamma})}{\left( \frac{1}{\gamma} \right)^{k+1}} \right]^2 \quad (7)$$

The above equation can be evaluated from [2, Eq. (8.352)]

$$F_{YSC}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{k! \left[ 1 - e^{-\frac{\gamma}{\gamma}} \left( \sum_{m=0}^k \frac{\left( \frac{\gamma}{\gamma} \right)^m}{m!} \right) \right]}{\left( \frac{1}{\gamma} \right)^{k+1}} \right]^2 \quad (8)$$

PDF of SC is  $f_{YSC}(\gamma) = \frac{dF_{YSC}(\gamma)}{d\gamma}$

$$f_{YSC}(\gamma) = \frac{d}{d\gamma} \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{k! \left[ 1 - e^{-\frac{\gamma}{\gamma}} \left( \sum_{m=0}^k \frac{\left( \frac{\gamma}{\gamma} \right)^m}{m!} \right) \right]}{\left( \frac{1}{\gamma} \right)^{k+1}} \right]^2$$

$$f_{YSC}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^4 \frac{(k!)^2}{\left( \frac{1}{\gamma} \right)^{(k+1)^2}} \frac{d}{d\gamma} \left[ 1 - e^{-\frac{\gamma}{\gamma}} \left( \sum_{m=0}^k \frac{\left( \frac{\gamma}{\gamma} \right)^m}{m!} \right) \right] \right]$$

$$f_{YSC}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^4 \frac{(k!)^2}{\left( \frac{1}{\gamma} \right)^{(k+1)^2}} 2 \left[ 1 - e^{-\frac{\gamma}{\gamma}} \left( \sum_{m=0}^k \frac{\left( \frac{\gamma}{\gamma} \right)^m}{m!} \right) \right] \right]$$

$$f_{YSC}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^4 \frac{(k!)^2}{\left( \frac{1}{\gamma} \right)^{(k+1)^2}} 2 \left[ 1 - \left[ \sum_{m=0}^k \frac{(\gamma)^{-m}}{m!} e^{-\frac{1}{\gamma}} \frac{d}{d\gamma} (e^{-\gamma} \gamma^m) \right] \right] \right]$$

$$f_{YSC}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^4 \frac{(k!)^2}{\left( \frac{1}{\gamma} \right)^{(k+1)^2}} 2 \left[ 1 - \left[ \sum_{m=0}^k \frac{(\gamma)^{-m}}{m!} e^{-\frac{1}{\gamma}} [m e^{-\gamma} \gamma^{m-1} - \gamma^m e^{-\gamma}] \right] \right] \right]$$

$$f_{YSC}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^4 \frac{(k!)^2}{\left( \frac{1}{\gamma} \right)^{(k+1)^2}} 2 \left[ 1 - \left[ \sum_{m=0}^k \frac{(\gamma)^{-m}}{m!} e^{-\frac{1}{\gamma}} e^{-\gamma} \gamma^{m-1} (m - \gamma) \right] \right] \right]$$

$$f_{YSC}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^4 \frac{(k!)^2}{\left( \frac{1}{\gamma} \right)^{(k+1)^2}} \left[ 2 \left( 1 - e^{-\frac{\gamma}{\gamma}} \left[ \sum_{m=0}^k \frac{\left( \frac{\gamma}{\gamma} \right)^m}{m!} \left[ \sum_{m=0}^k \frac{(\gamma)^{-m}}{m!} e^{-\frac{1}{\gamma}} e^{-\gamma} \gamma^{m-1} (m - \gamma) \right] \right) \right] \right] \right] \quad (9)$$

### 3.2 Maximal Ratio Combining (MRC)

This is the most complex scheme in which all branches are optimally combined at the receiver. MRC requires scaling and co-phasing of individual branch. In this all the signals are weighted according to their individual signal voltage to noise power ratios and then summed. Thus MRC produces an output SNR, which is equal to the sum of the individual SNRs. Best statistical reduction of fading is achieved by this method. SNR of MRC is calculated using convolutional properties.

#### Convolution Properties

The convolution of  $f_{\gamma}(\gamma)$  and  $f_{\gamma}(\gamma)$  is written  $f_{\gamma}(\gamma) \oplus f_{\gamma}(\gamma)$ , using convolutional operator  $\oplus$ . It is defined as the integral of the product of the two functions after one is reversed and shifted. As such, it is a particular kind of integral transform:

$$f_{\text{YMRC}}(\gamma) = f_{\gamma}(\gamma) \oplus f_{\gamma}(\gamma)$$

$$f_{\text{YMRC}}(\gamma) = \int_{-\infty}^{\infty} f_{\gamma}(\tau) f_{\gamma}(\gamma - \tau) d\tau$$

For functions  $(f_{\gamma}(\gamma), f_{\gamma}(\gamma))$  supported on only  $0 \rightarrow \infty$  (i.e., zero for negative arguments), the integration limits can be truncated, resulting in

$$f_{\text{YMRC}}(\gamma) = \int_0^{\gamma} f_{\gamma}(\tau) f_{\gamma}(\gamma - \tau) d\tau$$

PDF of MRC is  $f_{\text{YMRC}}(\gamma) = \int_0^{\gamma} f_{\gamma}(\tau) f_{\gamma}(\gamma - \tau) d\tau$

$$f_{\gamma}(\tau) = \left(\frac{1}{\bar{\gamma}}\right) e^{-\frac{(\tau+s^2)}{\bar{\gamma}}} \sum_{j=0}^{\infty} \frac{1}{\Gamma(j+1)j!} \left(\frac{\sqrt{\tau}s}{\bar{\gamma}}\right)^{2j} \quad (10)$$

$$f_{\gamma}(\gamma - \tau) = \left(\frac{1}{\bar{\gamma}}\right) e^{-\frac{(\gamma-\tau+s^2)}{\bar{\gamma}}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\sqrt{\gamma-\tau}s}{\bar{\gamma}}\right)^{2k} \quad (11)$$

PDF of MRC can be calculated using Eq. (10)&(11)

$$f_{\text{YMRC}}(\gamma) = \int_0^{\gamma} f_{\gamma}(\tau) f_{\gamma}(\gamma - \tau) d\tau$$

$$f_{\text{YMRC}}(\gamma) = \int_0^{\gamma} \left[ \frac{1}{\bar{\gamma}} e^{-\frac{(\tau+s^2)}{\bar{\gamma}}} \sum_{j=0}^{\infty} \frac{1}{\Gamma(j+1)j!} \left(\frac{\sqrt{\tau}s}{\bar{\gamma}}\right)^{2j} \right] \left[ \frac{1}{\bar{\gamma}} e^{-\frac{(\gamma-\tau+s^2)}{\bar{\gamma}}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\sqrt{\gamma-\tau}s}{\bar{\gamma}}\right)^{2k} \right] d\tau$$

$$f_{\text{YMRC}}(\gamma) = \int_0^{\gamma} \left[ \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \left(\frac{\sqrt{\tau}s}{\bar{\gamma}}\right)^{2j} \left(\frac{\sqrt{\gamma-\tau}s}{\bar{\gamma}}\right)^{2k} \right] d\tau$$

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \int_0^{\gamma} \left(\frac{\sqrt{\tau}s}{\bar{\gamma}}\right)^{2j} \left(\frac{\sqrt{\gamma-\tau}s}{\bar{\gamma}}\right)^{2k} d\tau$$

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \int_0^{\gamma} \left(\frac{\tau s^2}{\bar{\gamma}^2}\right)^j \left(\frac{(\gamma-\tau)s^2}{\bar{\gamma}^2}\right)^k d\tau$$

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \left(\frac{s}{\bar{\gamma}}\right)^{2j} \left(\frac{s}{\bar{\gamma}}\right)^{2k} \int_0^{\gamma} (\tau)^j (\gamma - \tau)^k d\tau$$

The above equation can be calculated from [2, Eq.1.111]

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \left(\frac{s}{\bar{\gamma}}\right)^{2j} \left(\frac{s}{\bar{\gamma}}\right)^{2k} \int_0^{\gamma} (\tau)^j \sum_{i=0}^n \binom{n}{i} (-\tau)^i (\gamma)^{n-i} d\tau$$

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \left(\frac{s}{\bar{\gamma}}\right)^{2j} \left(\frac{s}{\bar{\gamma}}\right)^{2k} \sum_{i=0}^n \binom{n}{i} (-1)^i (\gamma)^{n-i} \int_0^{\gamma} (\tau)^j (\tau)^i d\tau$$

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \left(\frac{s}{\bar{\gamma}}\right)^{2j} \left(\frac{s}{\bar{\gamma}}\right)^{2k} \sum_{i=0}^n \binom{n}{i} (-1)^i (\gamma)^{n-i} \int_0^{\gamma} (\tau)^{i+j} d\tau$$

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \left(\frac{s}{\bar{\gamma}}\right)^{2j} \left(\frac{s}{\bar{\gamma}}\right)^{2k} \sum_{i=0}^n \binom{n}{i} (-1)^i (\gamma)^{n-i} \frac{\gamma^{i+j+1}}{i+j+1}$$

$$f_{\text{YMRC}}(\gamma) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left(\frac{1}{\Gamma(j+1)j!}\right) \left(\frac{1}{\Gamma(k+1)k!}\right) \left(\frac{s}{\bar{\gamma}}\right)^{2j} \left(\frac{s}{\bar{\gamma}}\right)^{2k} \sum_{i=0}^n \binom{n}{i} (-1)^i \frac{\gamma^{n+j+1}}{i+j+1} \quad (12)$$

### Outage Probability $P_{\text{out}}$

$$P_{\text{out MRC}} = F_{\text{YMRC}}(u) = \int_0^u f_{\text{YMRC}}(\gamma) d\gamma$$

$$F_{Y_{MRC}}(u) = \int_0^u \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{1}{\Gamma(j+1)j!} \right) \left( \frac{1}{\Gamma(k+1)k!} \right) \left( \frac{s}{\bar{\gamma}} \right)^{2j} \left( \frac{s}{\bar{\gamma}} \right)^{2k} \sum_{i=0}^n \binom{n}{i} (-1)^i \frac{\gamma^{n+j+1}}{i+j+1} d\gamma$$

$$F_{Y_{MRC}}(u) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{1}{\Gamma(j+1)j!} \right) \left( \frac{1}{\Gamma(k+1)k!} \right) \left( \frac{s}{\bar{\gamma}} \right)^{2j} \left( \frac{s}{\bar{\gamma}} \right)^{2k} \sum_{i=0}^n \binom{n}{i} (-1)^i \int_0^u \frac{\gamma^{n+j+1}}{i+j+1} d\gamma$$

$$F_{Y_{MRC}}(u) = \frac{1}{\bar{\gamma}} e^{-(\gamma+2s^2)} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{1}{\Gamma(j+1)j!} \right) \left( \frac{1}{\Gamma(k+1)k!} \right) \left( \frac{s}{\bar{\gamma}} \right)^{2j} \left( \frac{s}{\bar{\gamma}} \right)^{2k} \sum_{i=0}^n \binom{n}{i} (-1)^i \frac{u^{n+j+2}}{(n+j+2)(i+j+1)} \quad (13)$$

Where  $\gamma$  is instantaneous SNR and  $\bar{\gamma}$  is average SNR

$$\gamma = |\alpha|^2$$

$k, i$  = running variable

### 3.3 Equal Gain Combining (EGC)

A variant of MRC is equal gain combining (EGC), where signals from each branch are co-phased and their weights have equal magnitude. This method also has possibility like MRC of producing an acceptable output signal from a number of unacceptable input signals. SNR improvement of EGC is better than selection combining but not better than MRC. SNR of EGC is given as

$$\gamma_{EGC} = \frac{(\alpha_1 + \alpha_2)^2}{2}$$

$$\gamma_{EGC} = \frac{\alpha^2}{2}$$

Here  $\alpha = \alpha_1 + \alpha_2$

$$f_{\alpha}(\alpha) = \int_0^{\alpha} f_{\alpha_1}(\tau) f_{\alpha_2}(\alpha - \tau) d\tau$$

Probability density function of envelop

$$f_{\alpha}(x) = \frac{x}{\sigma^2} e^{-\frac{(x^2+s^2)}{2\sigma^2}} I_0\left(\frac{xs}{\sigma^2}\right) \quad x \geq 0, \quad (14)$$

Where

$I_0(\cdot)$  is zero order modified Bessel function of the first kind

$$I_0\left(\frac{xs}{\sigma^2}\right) = \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{xs}{2\sigma^2}\right)^{2k} \text{ is evaluated from [2, Eq. (8.445)]}$$

By putting the value of  $I_0(\cdot)$  in Eq. 14 we get

$$f_{\alpha}(x) = \frac{x}{\sigma^2} e^{-\frac{(x^2+s^2)}{2\sigma^2}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{xs}{2\sigma^2}\right)^{2k} \quad (15)$$

$$f_{\alpha_1}(\tau) = \frac{\tau}{\sigma^2} e^{-\frac{(\tau^2+s^2)}{2\sigma^2}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\tau s}{2\sigma^2}\right)^{2k} \quad (16)$$

$$f_{\alpha_2}(\alpha - \tau) = \frac{(\alpha - \tau)}{\sigma^2} e^{-\frac{((\alpha - \tau)^2 + s^2)}{2\sigma^2}} \sum_{j=0}^{\infty} \frac{1}{\Gamma(j+1)j!} \left(\frac{(\alpha - \tau)s}{2\sigma^2}\right)^{2j} \quad (17)$$

Where  $\alpha_1$  is fading envelop seen by antenna 1 &  $\alpha_2$  is fading envelop seen by antenna 2

Using Eq. 16 and 17 we calculate the PDF of envelop for EGC

$$f_{\alpha}(\alpha) = \int_0^{\alpha} f_{\alpha_1}(\tau) f_{\alpha_2}(\alpha - \tau) d\tau$$

$$\begin{aligned}
 f_{\alpha}(\alpha) &= \int_0^{\alpha} \frac{\tau}{\sigma^2} e^{-\frac{(\tau^2+s^2)}{2\sigma^2}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\tau s}{2\sigma^2}\right)^{2k} \frac{(\alpha-\tau)}{\sigma^2} e^{-\frac{((\alpha-\tau)^2+s^2)}{2\sigma^2}} \sum_{j=0}^{\infty} \frac{1}{\Gamma(j+1)j!} \left(\frac{(\alpha-\tau)s}{2\sigma^2}\right)^{2j} d\tau \\
 f_{\alpha}(\alpha) &= \int_0^{\alpha} \frac{(\alpha\tau - \tau^2)}{\sigma^4} e^{-\frac{(\tau^2+s^2+(\alpha-\tau)^2+s^2)}{2\sigma^2}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\tau s}{2\sigma^2}\right)^{2k} \frac{1}{\Gamma(j+1)j!} \left(\frac{(\alpha-\tau)s}{2\sigma^2}\right)^{2j} d\tau \\
 f_{\alpha}(\alpha) &= \int_0^{\alpha} \frac{(\alpha\tau - \tau^2)}{\sigma^4} e^{-\frac{(2\tau^2+2s^2+\alpha^2-2\alpha\tau)}{2\sigma^2}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\tau s}{2\sigma^2}\right)^{2k} \frac{1}{\Gamma(j+1)j!} \left(\frac{\alpha s - \tau s}{2\sigma^2}\right)^{2j} d\tau \\
 f_{\alpha}(\alpha) &= \int_0^{\alpha} \frac{(\alpha\tau - \tau^2)}{\sigma^4} e^{-\frac{(\alpha^2)}{2\sigma^2}} e^{-\frac{(s^2)}{\sigma^2}} e^{-\frac{(\tau^2)}{\sigma^2}} e^{-\frac{(\alpha\tau)}{\sigma^2}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\tau s}{2\sigma^2}\right)^{2k} \frac{1}{\Gamma(j+1)j!} \left(\frac{\alpha s - \tau s}{2\sigma^2}\right)^{2j} d\tau \\
 f_{\alpha}(\alpha) &= \int_0^{\alpha} \frac{(\alpha\tau - \tau^2)}{\sigma^4} e^{-\frac{(\alpha^2)}{2\sigma^2}} e^{-\frac{(s^2)}{\sigma^2}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left(\frac{\tau s}{2\sigma^2}\right)^{2k} \frac{1}{\Gamma(j+1)j!} \left(\frac{\alpha s - \tau s}{2\sigma^2}\right)^{2j} d\tau \\
 f_{\alpha}(\alpha) &= \int_0^{\alpha} e^{-\frac{(\alpha^2)}{2\sigma^2}} e^{-\frac{(s^2)}{\sigma^2}} \sum_{k=0}^{\infty} \frac{(\alpha\tau - \tau^2)}{\sigma^4} \left(\frac{\alpha s - \tau s}{2\sigma^2}\right)^{2j} \frac{e^{-\frac{(\tau^2)}{\sigma^2}} e^{-\frac{(\alpha\tau)}{\sigma^2}}}{\Gamma(k+1)k!} \left(\frac{\tau s}{2\sigma^2}\right)^{2k} d\tau \\
 f_{\alpha}(\alpha) &= \int_0^{\alpha} e^{-\frac{(\alpha^2)}{2\sigma^2}} e^{-\frac{(s^2)}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \frac{\alpha^{2k} s^{2k} \tau^{2k+1} - \tau^{2k+2} s^{2k}}{\Gamma(j+1)j! \sigma^{4j+4} 2^{2j+1}} \frac{(\alpha s - \tau s)^{2j}}{\Gamma(k+1)k!} e^{-\frac{(\tau^2)}{\sigma^2}} e^{-\frac{(\alpha\tau)}{\sigma^2}} d\tau
 \end{aligned}$$

The above equation can be calculated from [2, Eq. (1.211)]

$$\begin{aligned}
 f_{\alpha}(\alpha) &= \int_0^{\alpha} e^{-\frac{(\alpha^2)}{2\sigma^2}} e^{-\frac{(s^2)}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \frac{\alpha^{2k} s^{2k} \tau^{2k+1} - \tau^{2k+2} s^{2k}}{\Gamma(j+1)j! \sigma^{4j+4} 2^{2j+1}} \sum_{a=0}^{2j} \binom{2j}{a} (-\tau s)^a \alpha s^{2j-a} \\
 &\quad \times \sum_{b=0}^{\infty} \frac{\left(-\frac{\tau^2}{\sigma^2}\right)^b}{b!} \sum_{c=0}^{\infty} \frac{\left(-\frac{\alpha\tau}{\sigma^2}\right)^c}{c!} \frac{1}{\Gamma(k+1)k!} d\tau \\
 f_{\alpha \text{EGC}}(\alpha) &= \int_0^{\alpha} e^{-\frac{(\alpha^2)}{2\sigma^2}} e^{-\frac{(s^2)}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \frac{\alpha^{2k} s^{2k} \tau^{2k+1} - \tau^{2k+2} s^{2k}}{\Gamma(j+1)j! \sigma^{4j+4} 2^{2j+1}} \binom{2j}{a} (-1)^a (\tau s)^a \\
 &\quad \times (-1)^b \left(\frac{\tau^2}{\sigma^2}\right)^b (-1)^c \left(\frac{\alpha\tau}{\sigma^2}\right)^c \frac{1}{\Gamma(k+1)k!} d\tau
 \end{aligned}$$

$$\begin{aligned}
 f_{\alpha}(\alpha) &= \int_0^{\alpha} e^{-\frac{(\alpha^2)}{2\sigma^2}} e^{-\frac{(s^2)}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \binom{2j}{a} \frac{1}{\Gamma(j+1)j! \sigma^{4j+4} 2^{2j+1} \Gamma(k+1)k!} \times (-1)^{a+b+c} (\alpha^{2k} s^{2k} \tau^{2k+1} \\
 &\quad - s^{2k} \tau^{2k+2}) \tau^{a+2b+c} s^a \frac{1}{\sigma^{2(b+c)} b! c!} d\tau
 \end{aligned}$$

$$f_{\alpha}(\alpha)$$

$$= \int_0^{\alpha} e^{-\left(\frac{\alpha^2}{2\sigma^2}\right)} e^{-\left(\frac{s^2}{\sigma^2}\right)} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \binom{2j}{a} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{2j+1}} \frac{1}{\Gamma(k+1)k!} \times (-1)^{a+b+c} (\alpha^{2k} s^{2k+a} \tau^{2k+a+2b+c+2}) \frac{\alpha^c}{\sigma^{2(b+c)} b! c!} d\tau$$

$$f_{\alpha}(\alpha)$$

$$= \int_0^{\alpha} e^{-\left(\frac{\alpha^2}{2\sigma^2}\right)} e^{-\left(\frac{s^2}{\sigma^2}\right)} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \binom{2j}{a} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{2j+1}} \frac{1}{\Gamma(k+1)k!} \times (-1)^{a+b+c} \frac{\alpha^c}{\sigma^{2(b+c)} b! c!} (\alpha^{2k} s^{2k+a} \tau^{2k+a+2b+c+2}) d\tau$$

$$f_{\alpha}(\alpha) = e^{-\left(\frac{\alpha^2}{2\sigma^2}\right)} e^{-\left(\frac{s^2}{\sigma^2}\right)} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \binom{2j}{a} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{2j+1}} \frac{1}{\Gamma(k+1)k!} \times (-1)^{a+b+c} \frac{\alpha^c}{\sigma^{2(b+c)} b! c!} \left( \alpha^{2k} s^{2k+a} \frac{\alpha^{2k+a+2b+c+2}}{2k+a+2b+c+2} - s^{2k+a} \frac{\alpha^{2k+a+2b+c+3}}{2k+a+2b+c+3} \right)$$

$$f_{\alpha}(\alpha) = e^{-\left(\frac{\alpha^2}{2\sigma^2}\right)} e^{-\left(\frac{s^2}{\sigma^2}\right)} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{2j+1}} \frac{1}{\Gamma(k+1)k!} \binom{2j}{a} \times (-1)^{a+b+c} \frac{s^{2k+a}}{\sigma^{2(b+c)} b! c!} \left( \frac{\alpha^{4k+a+2b+2c+2}}{2k+a+2b+c+2} - \frac{\alpha^{2k+a+2b+2c+3}}{2k+a+2b+c+3} \right) \quad (18)$$

### PDF of SNR is

By applying the value of  $\alpha = 2\sqrt{\gamma}$  in equation 18 we get

$$f_{\gamma_{EGC}}(\gamma) = \frac{1}{\sqrt{\gamma}} f_{\alpha_{EGC}}(2\sqrt{\gamma})$$

$$f_{\gamma_{EGC}}(\gamma) = e^{-\frac{4\gamma}{2\sigma^2}} e^{-\frac{s^2}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{j+1}} \binom{2j}{a} \times (-1)^{a+b+c} \frac{s^{2k+a}}{\sigma^{2(b+c)} b! c!} \left( \frac{2^{4k+a+2b+2c+2} \gamma^{2k+\frac{a}{2}+b+c+\frac{1}{2}}}{2k+a+2b+c+3} - \frac{2^{2k+a+2b+2c+3} \gamma^{k+\frac{a}{2}+b+c+1}}{2k+a+2b+c+3} \right) \quad (19)$$

### Outage Probability of Equal Gain Combining

$$P_{out} = F_{\gamma_{EGC}}(u) = \int_0^u f_{\gamma_{EGC}}(\gamma) d\gamma$$

$$F_{\gamma_{EGC}}(u) = \int_0^u e^{-\frac{4\gamma}{2\sigma^2}} e^{-\frac{s^2}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{j+1}} \binom{2j}{a} \times (-1)^{a+b+c} \frac{s^{2k+a}}{\sigma^{2(b+c)} b! c!} \left( \frac{2^{4k+a+2b+2c+2} \gamma^{2k+\frac{a}{2}+b+c+\frac{1}{2}}}{2k+a+2b+c+3} - \frac{2^{2k+a+2b+2c+3} \gamma^{k+\frac{a}{2}+b+c+1}}{2k+a+2b+c+3} \right) d\gamma$$

$$F_{\gamma_{EGC}}(u) = e^{-\frac{4\gamma}{2\sigma^2}} e^{-\frac{s^2}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{j+1}} \binom{2j}{a} \times (-1)^{a+b+c} \frac{s^{2k+a}}{\sigma^{2(b+c)} b! c!} \int_0^u \left( \frac{2^{4k+a+2b+2c+2} \gamma^{2k+\frac{a}{2}+b+c+\frac{1}{2}}}{2k+a+2b+c+3} - \frac{2^{2k+a+2b+2c+3} \gamma^{k+\frac{a}{2}+b+c+1}}{2k+a+2b+c+3} \right) d\gamma$$

$$F_{\gamma_{EGC}}(u) = e^{-\frac{4\gamma}{2\sigma^2}} e^{-\frac{s^2}{\sigma^2}} \sum_{k=0}^{\infty} \sum_{j=0}^{\infty} \sum_{a=0}^{2j} \sum_{b=0}^{\infty} \sum_{c=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \frac{1}{\Gamma(j+1)j!} \frac{1}{\sigma^{4j+4} 2^{j+1}} \binom{2j}{a} \times (-1)^{a+b+c} \frac{s^{2k+a}}{\sigma^{2(b+c)} b! c!} \int_0^u \left( \frac{2^{4k+a+2b+2c+2} \gamma^{2k+\frac{a}{2}+b+c+\frac{1}{2}}}{2k+a+2b+c+3} - \frac{2^{2k+a+2b+2c+3} \gamma^{k+\frac{a}{2}+b+c+1}}{2k+a+2b+c+3} \right) d\gamma$$



$$\times (-1)^{a+b+c} \frac{s^{2k+a}}{\sigma^{2(b+c)} b! c!} \left( \frac{2^{4k+a+2b+2c+2} u^{\frac{2k+\frac{a}{2}+b+c+\frac{3}{2}}{2k+\frac{a}{2}+b+c+\frac{3}{2}}}}{2k+a+2b+c+3} - \frac{2^{2k+a+2b+2c+3} u^{\frac{k+\frac{a}{2}+b+c+2}{k+\frac{a}{2}+b+c+2}}}{2k+a+2b+c+3} \right) \quad (20)$$

### 3.4 Switch and Stay Combining (SSC)

SSC further simplifies the complexities of SC. In this in place of continually connecting the diversity path with best quality, a particular diversity path is selected by the receiver till the quality of the path drops below a predetermined threshold. When it happens, then the receiver switches to another diversity path. This reduces the complexities relative to SC, because continuous and simultaneous monitoring of all diversity paths is not required. The CDF of SNR of SSC is given [5] as

Outage Probability  $P_{\text{out}}$  of SSC is evaluated from [5, Eq.(9.270)]

$$P_{Y_{\text{SSC}}}(\gamma) = \begin{cases} P_Y(\gamma_T) P_Y(\gamma) \gamma < \gamma_T \\ P_Y(\gamma) - P_Y(\gamma_T) + P_Y(\gamma) P_Y(\gamma_T) \gamma \geq \gamma_T \end{cases} \quad (21)$$

Where  $\gamma_T$  is switching threshold

$$P_Y(\gamma) = \frac{e^{-\frac{s^2}{\gamma}}}{\gamma} \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{\Gamma(k+1, \frac{\gamma}{\gamma})}{(\frac{1}{\gamma})^{k+1}} \right] \quad (22)$$

$$P_Y(\gamma_T) = \frac{e^{-\frac{s^2}{\gamma_T}}}{\gamma_T} \left[ \sum_{j=0}^{\infty} \left( \frac{s^j}{j! \gamma_T^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\gamma_T})}{(\frac{1}{\gamma_T})^{j+1}} \right] \quad (23)$$

The Outage Probability of SSC can be calculated from Eq. (22)&(23)

When  $\gamma < \gamma_T$

$$P_{Y_{\text{SSC}}}(\gamma) = P_Y(\gamma_T) P_Y(\gamma)$$

$$P_{Y_{\text{SSC}}}(\gamma) = \left[ \frac{e^{-\frac{s^2}{\gamma}}}{\gamma} \left[ \sum_{j=0}^{\infty} \left( \frac{s^j}{j! \gamma_T^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\gamma_T})}{(\frac{1}{\gamma_T})^{j+1}} \right] \right] \left[ \frac{e^{-\frac{s^2}{\gamma}}}{\gamma} \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{\Gamma(k+1, \frac{\gamma}{\gamma})}{(\frac{1}{\gamma})^{k+1}} \right] \right]$$

$$P_{Y_{\text{SSC}}}(\gamma) = \left[ \frac{e^{-\frac{s^2}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{j=0}^{\infty} \left( \frac{s^j}{j! \gamma_T^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\gamma_T})}{(\frac{1}{\gamma_T})^{j+1}} \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{\Gamma(k+1, \frac{\gamma}{\gamma})}{(\frac{1}{\gamma})^{k+1}} \right]$$

$$P_{Y_{\text{SSC}}}(\gamma) = \left[ \frac{e^{-\frac{(s^2)}{\gamma}}}{\gamma} \right]^2 \left[ \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{s^j}{j! \gamma_T^j} \right)^2 \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\gamma_T})}{(\frac{1}{\gamma_T})^{j+1}} \frac{\Gamma(k+1, \frac{\gamma}{\gamma})}{(\frac{1}{\gamma})^{k+1}} \right] \quad (24)$$

When  $\gamma \geq \gamma_T$

$$P_{Y_{\text{SSC}}}(\gamma) = P_Y(\gamma) - P_Y(\gamma_T) + P_Y(\gamma) P_Y(\gamma_T)$$

$$P_{Y_{\text{SSC}}}(\gamma) = \left[ \frac{e^{-\frac{s^2}{\gamma}}}{\gamma} \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \gamma^k} \right)^2 \frac{\Gamma(k+1, \frac{\gamma}{\gamma})}{(\frac{1}{\gamma})^{k+1}} \right] \right] - \left[ \frac{e^{-\frac{s^2}{\gamma_T}}}{\gamma_T} \left[ \sum_{j=0}^{\infty} \left( \frac{s^j}{j! \gamma_T^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\gamma_T})}{(\frac{1}{\gamma_T})^{j+1}} \right] \right]$$

$$\begin{aligned}
 & + \left[ \frac{e^{-\frac{(s^2)}{\bar{\gamma}}}}{\bar{\gamma}} \right]^2 \left[ \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{s^j}{j! \bar{\gamma}^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{j+1}} \left( \frac{s^k}{k! \bar{\gamma}^k} \right)^2 \frac{\Gamma(k+1, \frac{\gamma}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{k+1}} \right] \\
 P_{Y_{SSC}}(\gamma) &= \left[ \frac{e^{-\frac{(s^2)}{\bar{\gamma}}}}{\bar{\gamma}} \right] \left[ \sum_{k=0}^{\infty} \left( \frac{s^k}{k! \bar{\gamma}^k} \right)^2 \frac{\Gamma(k+1, \frac{\gamma}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{k+1}} \right] - \left[ \sum_{j=0}^{\infty} \left( \frac{s^j}{j! \bar{\gamma}^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{j+1}} \right] \\
 & + \left[ \frac{e^{-\frac{(s^2)}{\bar{\gamma}}}}{\bar{\gamma}} \right]^2 \left[ \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{s^j}{j! \bar{\gamma}^j} \right)^2 \left( \frac{s^k}{k! \bar{\gamma}^k} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{j+1}} \frac{\Gamma(k+1, \frac{\gamma}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{k+1}} \right] \quad (25)
 \end{aligned}$$

PDF of SSC is evaluated from [5, Eq.(9.274)]

$$P_{Y_{SSC}}(\gamma) = \frac{dP_{Y_{SSC}}(\gamma)}{d\gamma} = \begin{cases} P_Y(\gamma_T) p_Y(\gamma) \gamma < \gamma_T \\ 1 + P_Y(\gamma_T) p_Y(\gamma) \gamma \geq \gamma_T \end{cases} \quad (26)$$

Where PDF of SNR is

$$p_Y(\gamma) = \frac{1}{\bar{\gamma}} e^{-\frac{(\gamma+s^2)}{\bar{\gamma}}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left( \frac{\sqrt{\gamma} s}{\bar{\gamma}} \right)^{2k}$$

When  $\gamma < \gamma_T$

$$P_{Y_{SSC}}(\gamma) = P_Y(\gamma_T) p_Y(\gamma)$$

$$\begin{aligned}
 P_{Y_{SSC}}(\gamma) &= \left[ \frac{e^{-\frac{s^2}{\bar{\gamma}}}}{\bar{\gamma}} \sum_{j=0}^{\infty} \left( \frac{s^j}{j! \bar{\gamma}^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{j+1}} \right] \left[ \frac{1}{\bar{\gamma}} e^{-\frac{(\gamma+s^2)}{\bar{\gamma}}} \sum_{k=0}^{\infty} \frac{1}{\Gamma(k+1)k!} \left( \frac{\sqrt{\gamma} s}{\bar{\gamma}} \right)^{2k} \right] \\
 P_{Y_{SSC}}(\gamma) &= \frac{e^{-\frac{(\gamma+2s^2)}{\bar{\gamma}}}}{\bar{\gamma}^2} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{s^j}{j! \bar{\gamma}^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{j+1}} \frac{1}{\Gamma(k+1)k!} \left( \frac{\sqrt{\gamma} s}{\bar{\gamma}} \right)^{2k} \quad (27)
 \end{aligned}$$

When  $\gamma \geq \gamma_T$

$$P_{Y_{SSC}}(\gamma) = 1 + P_Y(\gamma_T) p_Y(\gamma)$$

$$P_{Y_{SSC}}(\gamma) = 1 + \frac{e^{-\frac{(\gamma+2s^2)}{\bar{\gamma}}}}{\bar{\gamma}^2} \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \left( \frac{s^j}{j! \bar{\gamma}^j} \right)^2 \frac{\Gamma(j+1, \frac{\gamma_T}{\bar{\gamma}})}{(\frac{1}{\bar{\gamma}})^{j+1}} \frac{1}{\Gamma(k+1)k!} \left( \frac{\sqrt{\gamma} s}{\bar{\gamma}} \right)^{2k} \quad (28)$$

#### IV SIMULATION RESULTS

The rician fading model is presented in fig.1 when the rice factor is 0 the channel works as rayleigh distribution and average SNR is equal to the received SNR. When the value of rice factor is 1 it works as Rician distribution and the average SNR increases. The outage probability is shown in fig.2 for K=0. It shows that at 8 dB SNR outage probability of SSC is  $10^{-1.7}$ , for SC is  $10^{-1.9}$ , for EGC is  $10^{-1.95}$  and for MRC  $10^{-2.0}$ , the outage performance improves from (from

SSC to MRC)  $10^{-1.7}$  to  $10^{-2.0}$ . Infig.3 the outage probability for  $K=1$  is shown by which it is determined that the outage probability improves to  $10^{-2.9}$  in case of MRC. Threshold of receiver is shown in fig.4 and 5 which indicates that receiver threshold is better for MRC and higher value of  $K$ . By these results we can conclude that MRC is giving better performance than other combining techniques. The Rice factor  $K$  plays a vital role for the performance of the channel when the value of  $K$  increases the performance improves.

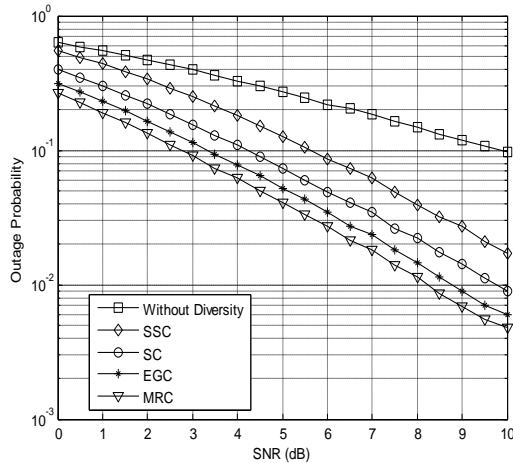


Fig.2. Outage Probability for  $K=0$

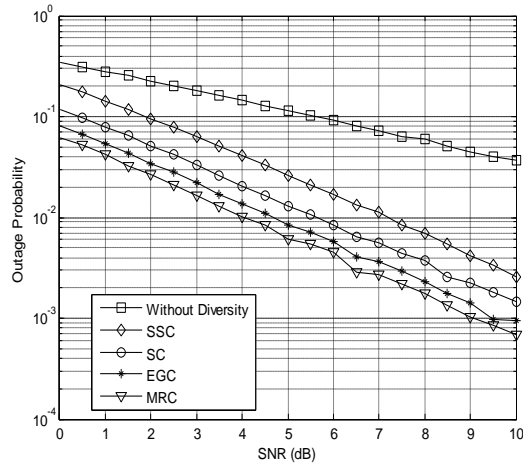


Fig.3. Outage Probability for  $K=1$

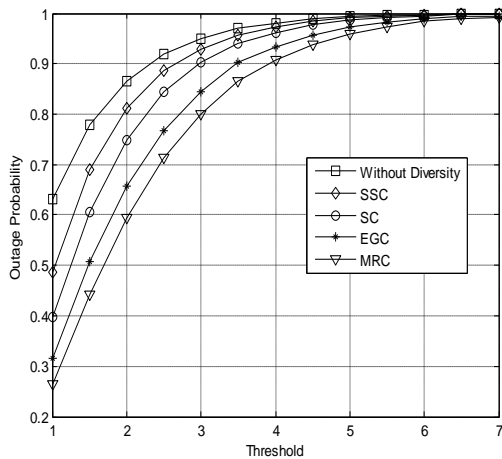
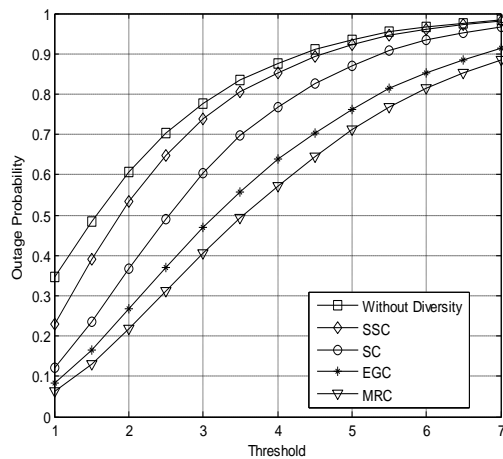


Fig.4 Threshold for Rician Fading Channel for  $K=0$  Fig.5 Threshold for Rician Fading Channel for  $K=1$



## V. CONCLUSION

In this paper probability distribution function, cumulative distribution function and outage probability of selection combining, maximal ratio combining, switch and stay combining, equal gain combining have been analysed. The simulation result shows that the outage performance of maximal ratio combining is better than all other combining techniques in both cases of the value of  $K$  is 0 as well as 1. By this analysis we can conclude that the combining technique and rice factor plays a vital role in Rician fading channel.

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## BIOGRAPHY



**Suryakant Pathak** received his MCA, M.Tech in Computer Science & Engineering and pursuing Ph.D. in Wireless Communication from BBD University Lucknow, India. He had worked in Indian Air Force, 1983-2003. Faculty at DOEEC Centre, Madan Mohan Malviya University of Technology campus Gorakhpur (U.P.), India, 2004-2006. Lecturer at Amity University Lucknow (U.P.), India, 2007-2008. Assistant Professor at Sagar Institute of Technology & Management Barabanki (U.P.), India, 2009-2013. At present he is Associate Professor & Head of Department Computer Science & Engineering Dr. K. N. Modi University Newai, Rajasthan, India. His research interests include almost all aspects of wireless communications with a special emphasis on MIMO systems, MIMO-OFDM, spread spectrum system, cooperative diversity schemes.



Himanshu Katiya received his B.E. degree in Electronics and Communication Engineering from M.I.T. Moradabad (M.J.P. Rohilkhand University, Bareilly) in 2001, M.Tech degree from Madan Mohan Malviya Engineering College, Gorakhpur (U.P. Technical University, Lucknow) in 2004 and Ph.D.

degree in the area of wireless communication at the Indian Institute of Technology, Guwahati, Assam India in 2011. From 2004–2005 he was lecturer of Electronics and Communication Engineering Dept. at SRM SCET, Bareilly, Uttar Pradesh, India and from 2005–2006 he was lecturer of Electronics and Communication Engineering Dept. at NIEC, Lucknow, and Uttar Pradesh, India. At present he is Associate Professor of the Electronics and Communication Engineering Dept. at BBDNIIT, Lucknow, and Uttar Pradesh, India. He was awarded from IETE research fellowship and he was project investigator (from September, 2009 to December, 2010) of an IETE sponsored research project. He has published over thirteen research papers in SCI journals, twelve research papers in non SCI journals and ten national/ international conference papers. His research interests include almost all aspects of wireless communications with a special emphasis on MIMO systems, MIMO-OFDM, spread spectrum system, channel modelling, infrastructure-based multi-hop and relay networks, cognitive radio communication, cooperative diversity schemes, adaptive array processing for Smart Antenna and Arduino based embedded system.